Global Carbon Fluxes Using Multioutput Gaussian Processes Regression and MODIS Products

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Abstract—The quantification of carbon fluxes (CFs) is crucial due to their role in the global carbon cycle having a direct impact on Earth's climate. In the last years, considerable efforts have been made to scale CFs from eddy covariance (EC) data to the globe. In this work, a data-driven approach that exploits a multioutput Gaussian processes regression algorithm (G-model) is proposed to jointly estimate gross primary production (GPP), terrestrial ecosystem respiration (TER), and net ecosystem exchange (NEE) at a global scale. The G-model not only provides an estimate of the CFs but also an uncertainty. Moreover, it derives the three fluxes jointly preserving their physical relationship. The predictors are selected from a set of the moderate-resolution imaging spectroradiometer (MODIS) products available on Google Earth engine. The performance of the model revealed high accuracies (R^2 reaching 0.82, 0.69, and 0.80 in the case of GPP, NEE, and TER, respectively), and low root mean square errors (1.55 g $m^{-2} d^{-1}$ in the case of GPP, 1.09 g m⁻² d⁻¹ for the NEE, and 1.14 g m⁻² d⁻¹ for TER) over the FLUXNET2015 data set at eight-day time scale. The GPP estimates provided by *G*-model outperformed the MOD17A2 product, and a state-of-the-art GPP product (PML_V2) without using meteorological forcing data sets. The results reported mean annual amounts of 133.7, 114.8, and 18.9 Pg yr⁻¹ for GPP, TER, and NEE, respectively, during the 2002-2023 period. The proposed approach paves the way for the development of multioutput strategies that preserve the physical relationships among CFs in upscaling processes.

Index Terms—Data-driven, gross primary production (GPP), machine learning (ML), multioutput Gaussian process regression, net ecosystem exchange (NEE), terrestrial ecosystem respiration (TER).

I. INTRODUCTION

C ARBON fluxes (CFs) are essential for understanding the Earth's carbon cycle and its influence on the climate system. The estimation of CFs is a measure of how much carbon

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is being exchanged between the surface and the atmosphere. These exchanges occur via natural processes such as photosynthesis, respiration, and decomposition. The three main CFs are the gross primary production (GPP), terrestrial ecosystem respiration (TER), and net ecosystem exchange (NEE). GPP stands for the total amount of carbon that is fixed by plants through photosynthesis, whereas TER refers to the total amount of carbon that is released into the atmosphere by the respiration of plants and other organisms present in the ecosystem. NEE is the difference between carbon uptake and carbon release by an ecosystem.

In-situ CFs are obtained from different networks of micrometeorological towers that use eddy covariance (EC) technique to estimate GPP from NEE and TER. Towers operate in many countries, across America, Europe, and Asia, through the networks AmeriFlux¹ [1], integrated carbon observation system $(ICOS)^2$ [2], AsiaFlux³ [3], or the FLUXNET global network⁴ [4]. These measurements are representative of the corresponding EC tower footprint, which may vary from hectometric to kilometric scales depending on the site characteristics. At regional and global scales CFs can be estimated using models that simulate the carbon cycle adding the influence of several factors such as vegetation types, soil properties, and climate conditions [5], [6], [7]. In addition, CFs can also be estimated from polar and geostationary remote sensing (RS) data such as the moderate resolution imaging spectroradiometer (MODIS) [8] or the spinning enhanced visible and infrared imager on board the meteosat second generation [9], respectively, which are based on production efficiency models (PEMs) based on Monteith logic. A simple yet effective method to estimate GPP from RS data is to employ vegetation indices or spectral band combinations as proxies of GPP [10], [11]. Similarly, solar-induced fluorescence can be used for inferring GPP [12], [13]. Other approaches link RS and meteorological data with EC measures by using machine learning (ML) methods [14], [15], [16], [17], [18], [19]. These data-driven techniques rely on the relationship between vegetation properties and CFs to build a single model (one per flux to be retrieved). The use of a single ML model for every predictive variable can lead to good independent performances but does not guarantee that the physical relationship among variables is respected, leading to nonrobust approaches. Some studies have

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¹[Online]. Available at: https://ameriflux.lbl.gov.

²[Online]. Available at: https://www.icos-cp.eu/.

³[Online]. Available at: http://www.asiaflux.net.

⁴[Online]. Available at: https://fluxnet.org/.

shown that the use of a unique multioutput model could improve the results when estimating biophysical parameters at the same time [20], [21]. However, this technique has not been tested on upscaling global CFs exchange yet, where the balance between carbon uptake through photosynthesis and carbon release through respiration leads to net ecosystem exchange GPP-TER = NEE. Therefore, it is necessary to develop novel machine learning algorithms that can preserve the physical consistency across CFs from local to global scales and intercompare the global estimates derived from various products.

In this framework, the objectives of this article are as follows.

- i) To develop a data-driven approach to estimate CFs at a global scale blending MODIS products and EC data.
- ii) To build a single multioutput model able to jointly estimate GPP, NEE, and TER in a robust way.
- iii) To validate the estimates against global in-situ data, and compare with two reference products (MOD17A2H and PML_V2).

The approach relies on the use of multitemporal MODIS products from Google Earth engine (GEE) as inputs in a multioutput Gaussian processes regression algorithm (\mathcal{G} -model), which exploits the covariance among CFs. The rest of this article is organized as follows. Section II describes in-situ and RS products, and the proposed multioutput algorithm. Section III shows the model performance, the explanatory power of the inputs, the derived global CFs maps, and a comparison with other products. Section IV discusses the results, and Section V concludes this article.

II. MATERIALS AND METHODS

A. Tower Data

The approach proposed in this study exploits data coming from the FLUXNET network. The FLUXNET2015 dataset was downloaded,⁵ which allowed the use of data coming from 211 sites globally distributed (see Supplementary Fig. SI). In particular, we used the dataset update dating from February 6, 2020. The distribution of the flux sites according to the International Geosphere-Biosphere Programme can be found in Supplementary Table SI. Daily GPP, NEE, and TER data were computed for every site from the average of the night- and day-time partitioning methods present in the FLUXNET2015 dataset [22]. The sign convention used in this study was such that CO_2 flux to the surface was positive so that NEE = GPP-TER, therefore positive values of NEE stand for vegetation carbon uptake, and negatives for release. The NEE quality flag was used to ensure the consistency between GPP, TER, and NEE, and similarly to other studies, we excluded daily data when more than 20% of the half-hourly data were based on gap-filling with low confidence [15], [23]. Finally, daily data were temporally aggregated (eight days) to match the RS products.

B. RS Products

The biosphere-atmosphere interactions related to photosynthesis are mainly driven by photosynthetically active radiation

 TABLE I

 MAIN FEATURES OF THE SELECTED MODIS-BASED PRODUCTS

_	GEE name	Spatial resolution	Temporal frequency	Reference	
-	MCD18C2.061	0.05°	3 h	[27]	
OIS	MCD43A4.061	500 m	Daily	[28]	
Predict	MOD11A2.061	1 km	8 days	[29]	
	MOD16A2.061	500 m	8 days	[30]	
	MCD15A2H.061	500 m	8 days	[31]	
or arison	MOD17A2H.061	500 m	8 days	[8]	
compa	PML_V2 0.1.7	500 m	8 days	[17]	

(PAR), as well as environmental and physiological variables, such as water availability, temperature, chlorophyll content, and leaf area index (LAI) [24]. On this basis, five MODIS-based RS products were used to extract the predictors for the regression method. Overall, eight predictors were selected, namely: PAR from the MCD18C2 product; day- and night-time land surface temperature (LST_D, LST_N) from the MOD11A2 product; evapotranspiration (ET) and potential ET (PET) from the MOD16A2 product, as well as the water stress factor (C_{ws}) computed as $C_{\rm ws} = \frac{\rm ET}{\rm PET}$; the kernel version of the normalized difference vegetation index (kNDVI) [25] computed from the MCD43A4 red and NIR bands; and LAI from the MCD15A2H product. KNDVI was selected since generalizes the NDVI and other indices such as NIRv [25], and its good performance to retrieve vegetation traits [26]. The GEE platform was used to download the selected MODIS products over the flux sites from 2002 (when starts the availability of all considered MODIS products) to 2014 (when ends the availability of the FLUXNET2015 data), thus fitting the coincident temporal range between the FLUXNET2015 dataset and the RS observations. The final value of the predictors that were assigned to every tower was the mean of the pixel values contained in a 500-m radius buffer weighted by their spatial extension. Other spatial buffers were initially tested but provided suboptimal results. The quality flag of every predictor was also used to discard bad-quality pixels. In the case of the MCD18C2 product, the values were integrated to obtain a daily-integrated PAR and aggregated to eight-day for temporal consistency among the rest of the predictors and EC data. Eventually, two reference GPP products (MOD17A2H, and PML V2) were downloaded from GEE with the same aforementioned processing to compare the GPP retrievals. Table I highlights the main features of the products.

C. Multioutput Gaussian Processes Regression (G-Model)

Gaussian processes (GPs) [32] are statistical techniques that offer a probabilistic approach for kernel regression methods in which the likelihood of unobserved latent functions is supposed to be modeled by a multivariate Gaussian prior. The approach makes use of a latent function of the inputs $f(\mathbf{x})$ (the RS predictors) plus constant Gaussian noise to estimate the outputs (the CF of interest) as $y = f(\mathbf{x}) + \varepsilon_n$, $\varepsilon_n \sim N(0, \sigma_n^2)$. The latent function is chosen to be a zero mean GP prior, or $(\mathbf{x}) \sim \text{GP}(0, k_{\theta})$, where

⁵[Online]. Available at: https://fluxnet.org/data/fluxnet2015-dataset/.

 k_{θ} is a covariance function parametrized by θ . The noise power is given by the hyperparameter σ_n^2 , and it is assumed that the noise ε_n follows a prior Gaussian distribution. Given the GP priors, samples drawn from f(x) at $\mathbf{x}_i = \{x_i^1, \ldots, x_i^B\}_{i=1}^N$ (*N* is the number of training samples, and *B* is the number of predictors) follow a joint multivariate zero-mean Gaussian distribution defined by the covariance matrix **K**, where $[\mathbf{K}]_{ij} = k_{\theta}(x_i, x_j)$.

The output of the model is computed from the predictive distribution described by the equations

$$p\left(\mathbf{y}_{*} | \mathbf{x}_{*}, D\right) = N\left(y_{*} | \mu_{\mathrm{GPR}*}, \sigma_{\mathrm{GPR}*}^{2}\right)$$
(1)

$$\mu_{\text{GPR}*} = \mathbf{k}_{*}^{\top} \left(\mathbf{K} + \sigma_{n}^{2} \mathbf{I}_{n} \right)^{-1} \mathbf{y} = \mathbf{k}_{*}^{\top} \alpha \qquad (2)$$

$$\sigma_{\text{GPR}*}^2 = \sigma^2 + k_{**} - \mathbf{k}_*^\top (\mathbf{K} + \sigma_n^2 \mathbf{I}_n)^{-1} \mathbf{k}_* \qquad (3)$$

where *D* stands for the training dataset (being \mathbf{x}_* and \mathbf{y}_* the input and output), $\mathbf{k}_* = [k(x_*, x_1), \dots, k(x_*, x_N)]$ is an $N \times 1$ vector and $k_{**} = k(x_*, x_*)$. The model provides a full posterior probability from which predictions ($\mu_{\text{GPR}*}$) and confidence estimates ($\sigma_{\text{GPR}*}^2$) can be computed. In the single output approach $\mathbf{y} \in \mathbb{R}^{N \times 1}$, the output of interest (e.g., GPP) is computed from the inputs (eight RS predictors, $\mathbf{x}_i \in \mathbb{R}^{N \times 8}$) as

$$\hat{y} = f(\mathbf{x}) = \sum_{i=1}^{N} \alpha_i k_\theta(\mathbf{x}_i, \mathbf{x}) + \alpha_0$$
(4)

where α_i is the weight assigned to each predictor \mathbf{x}_i , α_0 is the bias term, and k_{θ} is the covariance function evaluating the similarity between \mathbf{x} and all the *N* training spectra. For this task, the automatic relevance determination kernel was selected

$$\mathbf{K}\left(\mathbf{x}_{i},\mathbf{x}_{j}\right) = \nu \exp\left(-\sum_{b=1}^{B}\left(\mathbf{x}_{i}^{b}-\mathbf{x}_{j}^{b}\right)^{2}/2\sigma_{b}^{2}\right) + \sigma_{n}^{2}\delta_{ij} \quad (5)$$

where ν is a scaling factor, σ_n accounts for the standard deviation of the noise, *B* is the number of predictors (in our case, B = 8), and σ_b can be related to the spread of each predictor *b*. The model hyperparameters $\boldsymbol{\theta} = [\nu, \sigma_n, \sigma_1, \dots, \sigma_b]$ can be estimated by maximizing the marginal log-likelihood [32]

$$\log p\left(\mathbf{y} \left| \mathbf{x}_{i}, \theta \right. \right) = -\frac{1}{2} \mathbf{y}^{\top} \left(\mathbf{K} + \sigma_{n}^{2} \mathbf{I}_{n} \right)^{-1} \mathbf{y} - \frac{1}{2} \log \left| \mathbf{K} + \sigma_{n}^{2} \mathbf{I}_{n} \right| -\frac{N}{2} \log \left(2\pi \right).$$
(6)

The model can be adapted to deal with multioutput regression problems (\mathcal{G} -model) adapting the kernel hyperparameters for a unique kernel, which is able to deal with all the outputs. In the \mathcal{G} -model we need to optimize the hyperparameters considering that $\mathbf{y} \in \mathbb{R}^{N \times D}$ instead $\mathbf{y} \in \mathbb{R}^{N \times 1}$ where D is the total number of outputs (in our case, D = 3). Therefore, a global cost function summarizing all the cost functions (one per output) into a single one must be defined. Here, we use a global cost function C just as the squared sum of the standard model cost function of each output [32]

$$C = \sum_{d=1}^{D} logp(\mathbf{y}_d | \mathbf{x})^2$$

$$= \sum_{d=1}^{D} \left(-\frac{1}{2} \mathbf{y}_{d}^{\top} \boldsymbol{\alpha} - \sum_{j=1}^{N} \log \mathcal{L}_{jj} - \frac{N}{2} \log (2\pi) \right)^{2}$$
(7)

where L_{jj} is the Cholesky factorization of the covariance matrix for every output.

D. Algorithm Calibration and Performance

The \mathcal{G} -model was created with 14 562 samples obtained matching the EC data and the RS products at eight-day temporal resolution. The inputs underwent min-max scaling to transform them into the 0–1 range, as the predictors exhibited a substantial range of variation. Employing this technique is highly advisable to mitigate the risk of suboptimal performance by machine learning algorithms. A total of 70% of the samples were randomly selected for training and we assessed the results in the remaining 30% (test set unused during model training). Different random initializations were conducted and no significant changes in performance were obtained (not shown for brevity). Root mean square error (RMSE), mean error (ME), mean absolute error (MAE), and the coefficient of determination (R^2) were computed over the test set to evaluate the accuracy, bias, and the goodness-of-fit of the model.

In addition, the model behavior was tested by identifying the most relevant predictors for every estimated flux. This was achieved by means of a sensitivity analysis carried out through an added-noise permutation approach [33]. The inputs were perturbed with Gaussian white noise $N(0,\sigma^2)$, being σ^2 the 3% of the perturbed signal amplitude. The noise was added to a single predictor remaining the rest of the predictors unperturbed, and repeating this process for every predictor. The predictor relevance for every flux was computed as the difference between the accuracy obtained with no perturbation, and the obtained when the perturbation was applied. Finally, the results were normalized with respect to the most relevant predictor.

III. RESULTS

A. Assessment of the Model Performance Over Tower Data

Fig. 1 shows the direct validation obtained over the test set. The outputs of the model preserve the physical relationship among the three fluxes (see codistributions in Fig. SII). The accuracy metrics revealed good correspondence between estimations and in-situ measurements. It is worth mentioning the high R^2 values, which reach 0.82, 0.69, and 0.80 in the case of GPP, NEE, and TER respectively. RMSE was $1.55 \text{ g m}^{-2} \text{ d}^{-1}$ for GPP, 1.09 g m⁻² d⁻¹ for NEE, and 1.14 g m⁻² d⁻¹ for TER. The NEE and TER predictions revealed low and very similar MAE and ME (see Fig. 1) whereas GPP exhibits slightly higher MAE (1.04 g m⁻² d⁻¹) and lower ME (0.01 g m⁻² d⁻¹). The same scatterplots showing the biome of every data point are provided in supplementary material (see Fig. SIII) as well as the number of validation data for every year and biome (see Fig. SIV). In addition, Fig. 2(d) and (e) shows the performance of MOD17A2 and PML V2 products over the same test set. Note that these products do not provide NEE nor TER. The RMSEs and MAEs of MOD17A2 and PML V2 are



Fig. 1. CFs, (a) GPP, (b) NEE, and (c) TER obtained with the G-model over the EC validation set. ME, RMSE, and MAE units are g m⁻² d⁻¹. GPP estimates of (d) MOD15A2 and (e) PML_V2 GPP obtained over the EC validation set. ME, RMSE, and MAE units are g m⁻² d⁻¹.



Fig. 2. *G*-model predictor importance for (a) GPP, (b) NEE, and (c) TER.

2.44 g m⁻² d⁻¹ and 2.11 g m⁻² d⁻¹, and 1.66 g m⁻² d⁻¹ and 1.48 g m⁻² d⁻¹, respectively. The reported biases regarding EC data are -0.50 g m⁻² d⁻¹ and -0.31 g m⁻² d⁻¹ in the case of MOD17A2 and PML_V2, respectively.

The \mathcal{G} -model accuracy metrics were disaggregated per biome type as shown in Table II. Note that the results over deciduous needleleaf forests are obtained only over one available site and may not be representative of this biome. The best results in all metrics for GPP were obtained over WSA. In the case of NEE, the best accuracy metrics were observed over EBFs and deciduous needleleaf forests (for ME and R^2), and OSHs (for RMSE and MAE). The lowest errors for TER were reported over WSA, whereas the highest correlation was found over EBFs.

B. Sensitivity Analysis

In every permutation, the amount of noise added to the perturbed input was 3% of the signal amplitude. Other amounts ranging from 3% to 9% of the signal amplitude were also tested and provided the same results regarding inputs' relevance. Fig. 2 shows the relevance of every predictor for GPP, NEE, and TER retrievals. The relevance of every predictor was normalized with the value of the maximum relevance predictor in every case. The \mathcal{G} -model sensitivity analysis revealed the daily integrated PAR from the MCD18C2 product as the most influential input for both GPP and NEE estimates. LAI from the MOD15A2 product proved as the most relevant predictor for TER retrievals.

C. Mean Annual Fluxes

After the completion of the \mathcal{G} -model training phase, the model was effectively applied to derive global GPP, NEE, and TER maps. Figs. 3–5 show the mean annual predictions and uncertainties of every flux computed for the 2002–2023 period along with their latitudinal profiles. The spatial pattern of the GPP (see Fig. 3, top) and TER (see Fig. 4, top) shows that

	GPP			NEE			TER					
	ME	RMSE	MAE	R^2	ME	RMSE	MAE	R^2	ME	RMSE	MAE	R^2
CRO	0.62	2.56	1.93	0.57	0.37	1.26	0.91	0.77	0.32	1.26	0.96	0.58
CSH	0.61	1.18	0.92	0.69	0.14	0.64	0.53	0.31	0.14	0.62	0.51	0.73
DBF	0.23	1.77	1.26	0.81	0.29	1.23	0.94	0.80	0.20	1.13	0.89	0.60
DNF(*)	0.61	0.85	0.64	0.78	1.19	1.37	1.19	0.74	1.25	1.39	1.25	0.42
EBF	0.15	1.33	0.93	0.88	0.01	1.02	0.75	0.41	0.06	1.17	0.84	0.88
ENF	0.28	1.41	1.02	0.79	0.00	1.27	0.84	0.51	0.07	1.40	0.88	0.73
GRA	0.17	1.31	0.87	0.88	0.28	1.15	0.75	0.46	0.28	1.22	0.79	0.87
MF	0.25	1.52	1.18	0.74	0.02	0.94	0.75	0.75	0.17	1.03	0.84	0.52
OSH	0.07	0.73	0.53	0.74	0.07	0.48	0.34	0.23	0.16	0.52	0.37	0.72
SAV	0.15	0.94	0.63	0.83	0.04	0.57	0.43	0.76	0.05	0.57	0.44	0.87
WET	0.16	2.11	1.47	0.70	0.20	1.14	0.87	0.79	0.14	1.10	0.80	0.77
WSA	0.01	0.58	0.45	0.91	0.09	0.51	0.36	0.71	0.05	0.45	0.33	0.86

 TABLE II

 Accuracy Assessment Per Biome in the EC Validation Set

Best results are highlighted in bold. GPP, NEE; TER, ME, RMSE and MAE in g m⁻² d⁻¹. Codes: croplands (CRO), closed shrublands (CSH), deciduous broadleaf forest (DBF), deciduous needleleaf forests (DNF), evergreen broadleaf forests (EBF), evergreen needleleaf forests (ENF), grasslands (GRA), mixed forests (MF), open shrublands (OSH), savannas (SA), wetlands (WET), and woody savannas (WSA). (*) Only one site.



Fig. 3. Global average maps for annual GPP (top) and model uncertainty σ_{GPP} (bottom) computed for the 2002–2023 period.

the tropics, encompassing regions like Southeast Asia, central Africa, and South America, exhibit a notable high mean annual GPP. Conversely, high latitudes, western North America, and both central Asia and Australia are characterized by considerably lower mean annual GPP and TER. The NEE map (see Fig. 5, top) presents a similar spatial pattern. It is worth mentioning that

the spatial distribution of the mean annual uncertainty of every flux is slightly different. The σ_{GPP} (see Fig. 3, bottom) exhibits a relatively consistent latitudinal profile, hovering around a value of 0.2 kg m⁻² yr⁻¹. However, there is a zone between 62 and 65° N where the values are lower (see blueish colors in the north of the σ_{GPP} map). In the case of σ_{TER} and σ_{NEE}



Fig. 4. Global average maps for annual TER (top) and model uncertainty σ_{TER} (bottom) computed for the 2002–2023 period.

their spatial patterns are similar (see Figs. 4 and 5, bottom). Their latitudinal uncertainty values are broadly constant around 0.15–0.2 kg m⁻² yr⁻¹ with no major variations, however σ_{NEE} presents lower spatial variability (see Fig. 5, bottom).

The results reported global amounts of 133.7, 114.8, and 18.9 Pg yr⁻¹ for GPP, TER, and NEE, respectively, in the 2002–2023 period. Note that those values preserve the relationship among CFs (NEE = GPP – TER). In the case of MOD17A2 and PML_V2 products, we found a GPP global mean of 118.3 Pg yr⁻¹ and 143.6 Pg yr⁻¹, respectively. Note that in the case of the PML_V2 the period was 2002–2020 due to data availability.

D. Comparison With MOD17A2 and PML_V2 Products

The GPP estimates were compared with the provided ones by the MOD17A2, and the PML_V2 products available on GEE. The comparison was carried out at the site level (see Fig. 6). The R^2 values regarding the MOD17A2 and PML_V2 were 0.63 and 0.70, respectively. The *G*-model predictions present a positive bias (0.71 g m⁻² d⁻¹) regarding the MOD17A2, and a negative one with respect to the PML_V2 estimates (-0.46 g m⁻² d⁻¹). Table III shows the comparison per biome type. The best concordances regarding the MOD17A2 product have been found over OSHs in terms of RMSE, MAE, and R^2 , and over crops in terms of ME. Over closed shrublands have been found the highest agreement regarding the PML_V2 estimates, but the lowest R^2 has been observed regarding the MOD17A2. The lowest correspondences regarding the two reference datasets have been reported over crops.

Fig. 7 shows different maps of the mean annual GPP regarding the reference datasets ($\Delta_{GPP} = GPP_{\mathcal{G}model}$ –reference). In general, the \mathcal{G} -model agrees well with the PML_V2, however, a slight overestimation of the GPP estimates provided by the PML_V2 regarding the \mathcal{G} -model is reported at high latitudes (see Fig. 7, top). Fig. 7 (bottom) shows the MOD17A2 estimates systematically underestimate GPP regarding the \mathcal{G} -model (mainly in the tropics) except over some locations.

Finally, Fig. 8 shows the global GPP computed for every year during the 2002–2023 period in the case of the \mathcal{G} -model, the MOD17A2, and the PML_V2 product (2002–2020 period). Global GPP provided by the \mathcal{G} -model varies from 131.8 Pg C yr⁻¹ to 137.6 Pg yr⁻¹, whereas the MOD17A2 reported lower amounts in the same period (118.1 Pg yr⁻¹ to 119.9 Pg C yr⁻¹). The PML_V2 product showed a range of 128.5–133.9 Pg yr⁻¹ (from 2002 to 2020). The corresponding linear trends in global GPP reported by the \mathcal{G} -model, the PML_V2, and the MOD17A2 are 0.28±0.02, 0.21±0.03, and 0.09±0.02 Pg yr⁻¹ yr⁻¹, respectively.

IV. DISCUSSION

A. Model Performance and Explanatory Variables

The state-of-the-art machine learning approaches that provide the highest accuracies are based on a combination of RS observations and meteorological variables as inputs [18], [19],



Fig. 5. Global average maps for annual NEE (top) and model uncertainty σ_{NEE} (bottom) computed for the 2002–2023 period.



Fig. 6. GPP comparison between the proposed approach and (a) the MOD17A2 and (b) the PML_V2 products, respectively. ME, RMSE, and MAE units are g $m^{-2} d^{-1}$.

[20]. Nevertheless, this study exclusively utilized a set of RS predictors derived from MODIS data into a machine learning regression method, resulting in a remarkable level of concordance between GPP, NEE, and TER retrievals and in-situ data. The interactions between CO_2 fluxes and RS/meteo variables depend significantly on ecosystem types and state [34]. A main advantage of the proposed approach is its potential to obtain globally those fluxes without a priori knowledge of the pixel characteristics. The results of the accuracy assessment show an

outperformance of the \mathcal{G} -model regarding both MOD17A2, and PML_V2 GPP products. The highest correlations for GPP and TER were obtained over WSA, SA, grasslands, (in line with the results of Joiner et al. [23]), and EBFs. In the case of NEE, the best correlation was obtained over the deciduous broadleaf forest, WET, and crops. Crops are the biome where the worst results were reported in the case of GPP and TER ($R^2 \leq 0.58$), however, the model performance in terms of correlation for NEE over this biome was good ($R^2 = 0.77$). Closed and OSHs are the

		Vs MOI	D17A2		Vs PML_V2				
	ME	RMSE	MAE	R^2	ME	RMSE	MAE	R^2	
CRO	0.84	2.51	1.91	0.35	-1.23	3.26	2.48	0.26	
CSH	0.82	1.48	1.05	0.26	-0.84	1.03	0.87	0.83	
DBF	0.20	1.90	1.39	0.67	-0.70	1.72	1.21	0.79	
DNF(*)	-0.80	1.38	1.50	0.81	-0.98	1.16	0.98	0.88	
EBF	0.21	2.11	1.60	0.64	-0.17	1.55	1.17	0.77	
ENF	0.44	1.98	1.46	0.49	-0.40	1.71	1.17	0.61	
GRA	0.28	1.88	1.31	0.67	-0.29	1.53	1.03	0.78	
MF	0.36	2.17	1.55	0.51	-0.56	2.29	1.57	0.52	
OSH	-0.72	1.22	0.90	0.72	-0.68	1.07	0.78	0.80	
SAV	0.21	1.24	0.92	0.56	0.07	1.00	0.70	0.68	
WET	0.38	2.18	1.89	0.49	-0.48	2.17	1.59	0.54	
WSA	-0.22	1.27	0.93	0.59	-0.21	0.78	0.57	0.81	

TABLE III GPP ACCURACY METRICS PER BIOME WITH REGARD TO THE MOD17A2 AND PML_V2 PRODUCTS

Best results are highlighted in bold. GPP, NEE; TER, ME, RMSE and MAE in g m⁻² d⁻¹. Codes: croplands (CRO), closed shrublands (CSH), deciduous broadleaf forest (DBF), deciduous needleleaf forests (DNF), evergreen broadleaf forests (EBF), evergreen needleleaf forests (ENF), grasslands (GRA), mixed forests (MF), open shrublands (OSH), savannas (SA), wetlands (WET), and woody savannas (WSA). (*) Only one site



Fig. 7. Global difference maps for mean annual GPP regarding (top) PML_V2 product and (bottom) MOD17A2 for the 2002–2020 period.

biomes where the lowest correlations for NEE ($R^2 \le 0.31$) have been found.

The \mathcal{G} -model consistently predicted GPP, NEE, and TER while effectively maintaining their inherent relationship. This is an important feature of this work since the usual procedure in machine learning approaches does not exploit the covariance of the CFs when upscaling in-situ data to a global scale.

Tramontana et al. [35] developed an approach for partitioning CFs preserving the relationship among them using neural networks physically constrained, but only applied with in-situ data.

The selection of the RS predictive variables is based on the main drivers of photosynthesis, transpiration, and respiration processes. In this regard, MODIS products that are related to water, radiation, temperature, and ET were selected. The ranking of



Fig. 8. Trends for mean annual GPP obtained by the G-model and the MOD17A2 for the 2002–2023 period, and the PML_V2 for 2002 to 2020.

the explanatory variables highlighted PAR as the most influential predictor of GPP estimates as expected according to Monteith logic [36]. However, the vegetation index kNDVI has also a great influence, which suggests that it might be connected with the light use efficiency (LUE). In fact, Wu et al. [37] proposed a GPP approach incorporating vegetation indices for LUE and the fraction of absorbed PAR (FAPAR) estimation. In literature, different authors have shown the usefulness of vegetation indices as proxies of both LUE and FAPAR [38], [39].

The relevance analysis for TER estimates revealed LAI as the most important variable. LST_N, ET, PET, and C_{ws} also revealed as important factors for TER. In the case of NEE, the main explanatory variables were the two most relevant for GPP and TER, i.e., PAR and LAI. Globally, the temperature is an influencing factor on the three fluxes: the diurnal temperature slightly affects GPP (i.e., photosynthesis), probably through the LUE in case of thermal stress (similar to the role of ET in case of water stress), and night temperature affects respiration process and, to a lesser extent, the NEE. Respiration provides the energy for a plant to acquire nutrients and to produce and maintain biomass. The total plant respiration can be split into three functional components: growth respiration, maintenance respiration, and the respiratory cost of ion uptake [40]. The maintenance respiration is expected to be greater in productive ecosystems and then to be positively correlated with LAI, with rather high relevance in respiration estimation as shown in Fig. 3. Maintenance respiration depends also on environmental stresses: for a given ecosystem it increases with temperature and drought. However, plants from hot environments have lower respiration rates at a given temperature than plants from cold places [40]. In any case, temperature and water balance are both rather relevant variables in respiration processes.

B. Global Estimates

GPP estimates at the site level provided by the \mathcal{G} -model are highly correlated with both MOD17A2 and PML_V2 estimates ($R^2 = 0.63$ and $R^2 = 0.71$). One factor that explains the high correlation is the use of similar inputs in the approaches. The per biome type comparison found the best agreements over OSHs in the case of MOD17A2, and over closed shrublands in the case of PML_V2. It is worth mentioning that the worst consistencies between products were obtained over crops in both cases. The spatial distribution of the GPP is consistent with the MOD17A2, and PML_V2 products. The major concentrations of GPP and TER are located in the tropics and decrease to high and low latitudes. In particular, the highest vegetation density areas (associated with high carbon uptakes) are distributed around the equator due to the presence of the intertropical convergence zone, a band of cumulus convection with a high average annual precipitation [41]. The lowest values are found over arid regions (such as Sahara and Arabian deserts in the North hemisphere and Atacama and Kalahari in the South hemisphere), which are located in the subtropical high-pressure belts centered around $\pm 30^{\circ}$ latitude.

The \mathcal{G} -model provided a global GPP (133.7 Pg yr⁻¹) similar to the one reported by other studies. For instance, Zhang et al. [42] reported an average global annual GPP around 125 Pg yr⁻¹ for the VPM and FluxCom products. Joiner et al. [23] revealed an annual GPP value of around 140 Pg yr^{-1} . Jung et al. [18] found a mean global GPP of 111 Pg yr⁻¹ for FluxCom-RS, 120 Pg yr⁻¹ for FluxCOM-RS+METEO product, and a range of 83–172 Pg yr⁻¹ for TRENDY v7 models. Dong et al. [43] carried out an inter-comparison of global GPP annual mean provided by PEM models, which reported a range of variation of 125–165 Pg yr⁻¹. In addition, Badgley et al. [10] reported a range of 131–163 Pg yr⁻¹ using the NIRv as a proxy for GPP. Recently, Li et al. [44] reported 125.74 Pg yr⁻¹ as a mean value for the 1982–2019 period, and Tagesson et al. [45] found 121.8 Pg yr⁻¹ over the 1982–2015 period. The TER and NEE budgets found in this study (114.8 Pg yr⁻¹, and 18.9 Pg yr⁻¹, respectively) are closer to the ones reported by Zeng et al. [46] (TER ranging from 115 to 121 Pg yr⁻¹, and NEE ranging from 20 to 22 Pg yr⁻¹). In addition, Li et al. [44] found 109.3 Pg yr⁻¹ for TER, and 16.28 Pg yr⁻¹ for NEE, and Tagesson et al. [47] reported 105.6 Pg yr⁻¹ for TER in the 1982–2015 period. These values are lower than those found in our study. The range of variation of the values reported in the literature indicates a significant level of uncertainty persists across different approaches.

The 2002–2023 temporal analysis provided a GPP trend $(0.28\pm0.02 \text{ Pg yr}^{-1} \text{ yr}^{-1})$ similar to other studies. Tagesson et al., [45] found a trend of $0.27\pm0.02 \text{ Pg yr}^{-1} \text{ yr}^{-1}$ in the 1982–2015 period, Guo et al. [19] and Zeng et al. [46] reported 021 Pg yr^{-1} yr^{-1} and 0.49 Pg yr^{-1} yr^{-1} both computed during the 1999–2019 period. The inter-comparison conducted by Dong et al. [43] reported a range of global GPP trends of different products, which varies from -0.22 to 0.51 Pg yr^{-1} yr^{-1} over different periods.

C. Model Uncertainty

Although the results show a good level of agreement regarding in-situ data and other products, there exist sources of uncertainty that propagate into the model. The spatial mismatch between the footprint of the EC towers and the RS data spatial resolution is a source of uncertainty in CF retrievals. Uncertainties in the EC covariance data (modeled by the noise parameter σ_n^2), and in the RS-based products can propagate into CFs retrieval [48], [49], [50]. In addition, the different spatial resolutions of the model inputs may also affect the stability of the predictions [51].

The \mathcal{G} -model provides not only a joint estimate of GPP, NEE, and TER but also an uncertainty (or confidence interval) for the retrievals. Note that the uncertainty of the predictive mean estimate does not depend on the value of the predictive mean but accounts for the similarity of the new inputs (x_*) regarding the inputs trained in the model (x_i). Therefore, high uncertainties would be obtained in scenarios where the behavior of the inputs is not accounted for during model training. The latitudinal profiles of the uncertainties are quite constant and do not follow the latitudinal variation of the GPP, NEE, and TER. This decorrelation between predictions and uncertainties suggests that the uncertainties are not influenced by the value of the prediction.

V. CONCLUSION

This study developed a data-driven approach using a multioutput regression method that is able to predict GPP, TER, and NEE at the same time. The \mathcal{G} -model estimates the three fluxes jointly also conserving their physical relationship, which is one of the drawbacks of the classic machine learning algorithms. The predictors of the model came from five MODIS products, which resulted in eight explanatory variables. Among them, the PAR and kNDVI were identified as the most relevant inputs for GPP, whereas the LAI and LST_N were the most influential for TER retrievals. PAR and LAI were also the most important for the NEE. The model was well calibrated at tower level and successfully upscaled to globe. At tower level, the best results were obtained for GPP ($R^2 = 0.82$), followed by TER ($R^2 = 0.80$), and NEE $(R^2 = 0.69)$. The RMSE was 1.55 g m⁻² d⁻¹ in the case of GPP, 1.09 g m⁻² d⁻¹ for NEE, and 1.14 g m⁻² d⁻¹ for TER. These results, obtained from only MODIS-based inputs, outperformed MOD17A2H and PML_V2 products. At a global scale, the spatial distribution of the derived maps agreed with MOD17A2H and PML_V2. The results provide a mean annual GPP value of 133.7 Pg yr⁻¹ and a global trend of 0.28 Pg yr⁻¹ yr⁻¹ for the 2002–2023 period that fall within the range reported by other works. However, further studies are needed to better understand the spatial and temporal patterns of global CFs that help to reduce uncertainties among approaches.

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